

Eco 5316 Time Series Econometrics  
Lecture 9 Seasonal Models

## Seasonal Models

- ▶ retail sales, labor force, unemployment rate, construction spending, housing starts and total miles traveled by air are just some examples of a large number of series that exhibit regular monthly or quarterly periodic pattern
- ▶ a lot of time series are officially published after performing seasonal adjustment that theoretically should remove seasonal component
- ▶ it is however not uncommon to encounter officially published seasonally adjusted data in which the seasonal pattern is still present

## Pure Seasonal Models

- ▶ simple pure seasonal AR model

$$y_t = \phi_s y_{t-s} + \varepsilon_t$$

ACF: spike at each multiple of  $s$  (lags 4, 8, 12, ... in case of quarterly data, lags 12, 24, 36, ... in case of monthly data)

PACF: single spike at lag  $s$

- ▶ simple pure seasonal MA model

$$y_t = \varepsilon_t + \theta_s \varepsilon_{t-s}$$

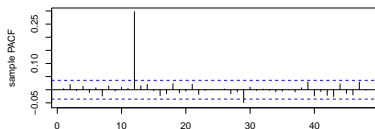
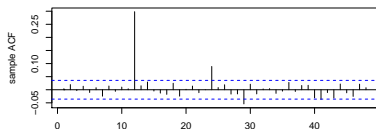
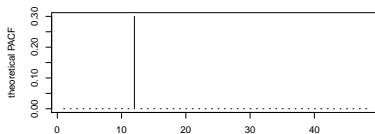
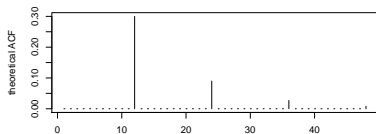
PACF: spike at each multiple of  $s$  (lags 4, 8, 12, ... in case of quarterly data, lags 12, 24, 36, ... in case of monthly data)

ACF: single spike at lag  $s$

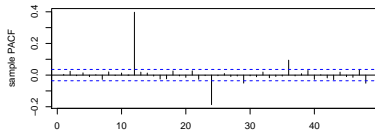
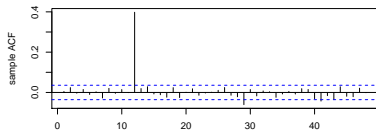
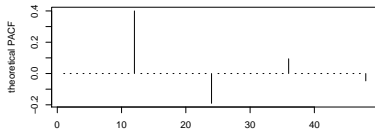
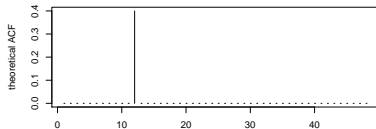
- ▶ in practice however most time series contain a seasonal AR or MA component as well as a regular AR or MA component

# Pure Seasonal Models

- ▶ example:  $s = 12$  and  $\phi_s = 0.3$



- ▶ example:  $s = 12$  and  $\theta_s = 0.5$



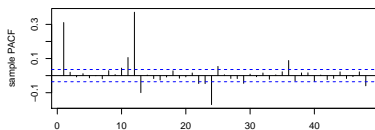
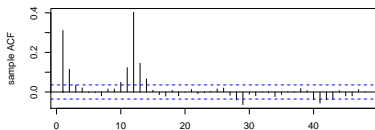
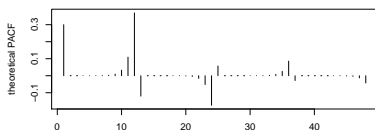
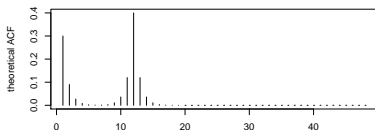
## Additive Seasonal AR model

- ▶ AR model with an additive seasonal MA component

$$(1 - \phi)y_t = (1 + \Theta L^s)\varepsilon_t$$

so that  $y_t = \phi y_{t-1} + \varepsilon_t + \Theta \varepsilon_{t-s}$

- ▶ if  $\phi > 0$ ,  $\Theta > 0$ 
  - ▶ ACF will exhibit exponential decay interrupted by a rise in correlation coefficients around lag  $s$  but no similar increase around  $2s$ ,  $3s$ , ...
  - ▶ PACF will be non-zero at lag 1 and unlike nonseasonal AR(1) model also show a pattern of increasing non-zero elements before *each* multiple of  $s$
- ▶ example:  $s = 12$  and  $\phi = 0.3$ ,  $\Theta = 0.5$



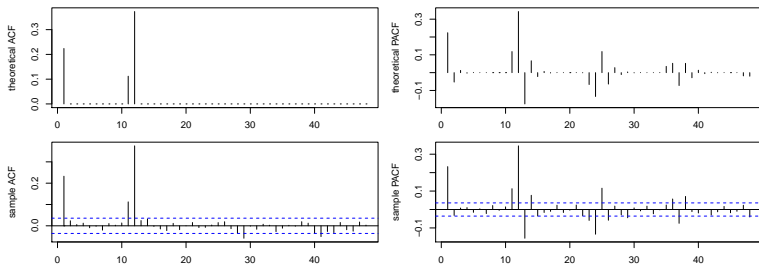
## Additive Seasonal MA model

- ▶ MA model with an additive seasonal MA component

$$y_t = (1 + \theta L + \Theta L^s) \varepsilon_t$$

so that  $y_t = \varepsilon_t + \theta \varepsilon_{t-1} + \Theta \varepsilon_{t-s}$

- ▶ ACF:  $\rho_1 \neq 0$ ,  $\rho_{s-1} \neq 0$ ,  $\rho_s \neq 0$ , all other elements zero
- ▶ PACF: larger non-zero elements around each multiple of  $s$
- ▶ example:  $s = 12$  and  $\theta = 0.3$ ,  $\Theta = 0.5$



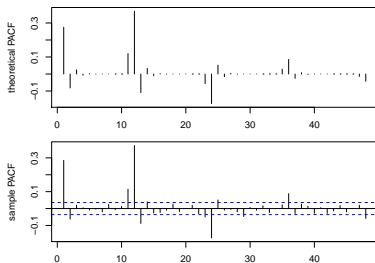
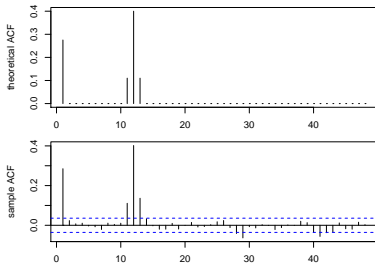
## Multiplicative Seasonal MA model

- ▶ Multiplicative seasonal MA model

$$y_t = (1 + \theta L)(1 + \Theta L^s)\varepsilon_t$$

so that  $y_t = \varepsilon_t + \theta\varepsilon_{t-1} + \Theta\varepsilon_{t-s} + \theta\Theta\varepsilon_{t-s-1}$

- ▶ ACF:  $\rho_1 \neq 0$ ,  $\rho_{s-1} \neq 0$ ,  $\rho_s \neq 0$ ,  $\rho_{s+1} \neq 0$ , all other elements zero
- ▶ PACF: larger non-zero elements around each multiple of  $s$
- ▶ compared to the additive model, multiplicative model allows for *interaction of regular and seasonal components*
- ▶ example:  $s = 12$  and  $\theta = 0.3$ ,  $\Theta = 0.5$



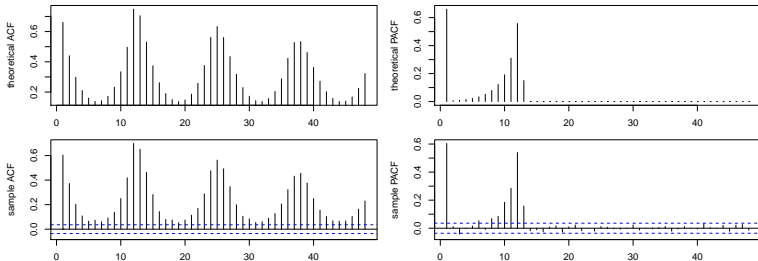
## Multiplicative Seasonal AR model

- ▶ multiplicative AR model with seasonal component

$$(1 - \phi L)(1 - \Phi L^s)y_t = \varepsilon_t$$

so that  $y_t = \phi y_{t-1} + \Phi y_{t-s} + \phi \Phi y_{t-s-1} + \varepsilon_t$

- ▶ if  $\phi > 0$ ,  $\Phi > 0$ 
  - ▶ ACF: exponential decay interrupted by increasing autocorrelations around *each* multiple of  $s$
  - ▶ PACF: large spikes at lag 1 and lag  $s$  and multiple smaller spikes between lag 2 and lag  $s+1$
- ▶ example:  $\phi_1 = 0.3$ ,  $\Phi_1 = 0.5$ .





## Multiplicative Seasonal AR model

- ▶ Q: why are lags 2 to 11 significant in PACF for  $(1-\phi L)(1-\Phi L^{12})y_t = \varepsilon_t$ , that is, if  $y_t$  is generated from

$$y_t = \phi y_{t-1} + \Phi y_{t-12} + \phi \Phi y_{t-13} + \varepsilon_t \quad (1)$$

- ▶ A: note that the above model implies that

$$y_{t-10} = \phi y_{t-11} + \Phi y_{t-22} + \phi \Phi y_{t-23} + \varepsilon_{t-10} \quad (2)$$

$$y_{t-11} = \phi y_{t-12} + \Phi y_{t-23} + \phi \Phi y_{t-24} + \varepsilon_{t-11} \quad (3)$$

and consider the OLS regressions to obtain PACFs for lags 10 to 14

$$(OLS_{10}) \quad y_t = \beta_0 + \beta_1 y_{t-1} + \dots + \beta_{10} y_{t-10} + e_t$$

$$(OLS_{11}) \quad y_t = \beta_0 + \beta_1 y_{t-1} + \dots + \beta_{10} y_{t-10} + \beta_{11} y_{t-11} + e_t$$

$$(OLS_{12}) \quad y_t = \beta_0 + \beta_1 y_{t-1} + \dots + \beta_{10} y_{t-10} + \beta_{11} y_{t-11} + \beta_{12} y_{t-12} + e_t$$

$$(OLS_{13}) \quad y_t = \beta_0 + \beta_1 y_{t-1} + \dots + \beta_{10} y_{t-10} + \beta_{11} y_{t-11} + \beta_{12} y_{t-12} + \beta_{13} y_{t-13} + e_t$$

- ▶  $\hat{\beta}_{13} \neq 0$  in  $OLS_{13}$ :  $y_t$  depends on  $y_{t-13}$  in (1) as long as  $\phi \neq 0$ ,  $\Phi \neq 0$
- ▶  $\hat{\beta}_{12} \neq 0$  in  $OLS_{12}$ :  $y_t$  depends on  $y_{t-12}$  in (1) as long as  $\Phi \neq 0$
- ▶  $\hat{\beta}_{11} \neq 0$  in  $OLS_{11}$ :  $y_t$  depends on  $y_{t-12}$  in (1),  $y_{t-11}$  depends on  $y_{t-12}$  in (2), there is omitted variable, since  $y_{t-12}$  does not appear in  $OLS_{11}$
- ▶  $\hat{\beta}_{10} \neq 0$  in  $OLS_{10}$ :  $y_t$  depends on  $y_{t-12}$  in (1),  $y_{t-10}$  depends on  $y_{t-11}$  in (2),  $y_{t-11}$  depends on  $y_{t-12}$  in (3), there is omitted variable bias,  $y_{t-12}$  does not appear in  $OLS_{10}$

## Seasonal Differencing

- ▶ we discussed that for economic data that is nonstationarity due to economic growth a common approach is to transform data using a logarithm and apply regular differencing

$$w_t = \Delta \log y_t$$

where  $\Delta = 1 - L$ , so that  $w_t = (1 - L) \log y_t$

- ▶ for economic data that is both nonstationarity due to economic growth and shows seasonal pattern the approach is to transform data using a logarithm and apply both regular and seasonal differencing

$$w_t = \Delta_s \Delta \log y_t$$

where  $\Delta = 1 - L$  and  $\Delta_s = 1 - L^s$ , so we have  $w_t = (1 - L^s)(1 - L) \log y_t$

- ▶ occasionally, when multiple unit roots are present, data has to be differenced more than once by applying  $\Delta^d = (1 - L)^d$  or  $\Delta_s^D = (1 - L^s)^D$

## General Multiplicative ARIMA model

- ▶ multiplicative models are written in the form  $\text{ARIMA}(p, d, q)(P, D, Q)_s$

$$(1 - \phi_1 L - \dots - \phi_p L^p)(1 - \Phi_1 L^s - \dots - \Phi_P L^{sP}) \Delta_s^D \Delta^d y_t \quad (4)$$

$$= (1 - \theta_1 L - \dots - \theta_q L^q)(1 - \Theta_1 L^s - \dots - \Theta_Q L^{sQ}) \varepsilon_t \quad (5)$$

- ▶ in practice  $\text{ARIMA}(1, 1, 0)(0, 1, 1)_s$  and  $\text{ARIMA}(0, 1, 1)(0, 1, 1)_s$  occur routinely

$$(1 - \phi_1 L) \Delta_s \Delta y_t = (1 - \Theta_1 L^s) \varepsilon_t$$

$$\Delta_s \Delta y_t = (1 - \theta_1 L)(1 - \Theta_1 L^s) \varepsilon_t$$

## Example: Johnson & Johnson quarterly earnings per share

```
# load necessary packages
library(readr)
library(dplyr)
library(tidyr)
library(purrr)
library(ggplot2)
library(ggfortify)
library(zoo)
library(timetk)
library(tibbletime)
library(lubridate)
library(forecast)
library(broom)
library(sweep)

# import the data on earnings per share for Johnson and Johnson
# then construct log, change, log-change, seasonal log change
tbl.wide.all <-
  read_table("http://faculty.chicagobooth.edu/ruey.tsay/teaching/fts3/q-jnj.txt", col_names = "y") %>%
  ts(start = c(1960,1), frequency = 4) %>%
  tk_tbl() %>%
  mutate(ly = log(y),
         dy = y - lag(y),
         dly1 = ly - lag(ly),
         dly4 = ly - lag(ly, 4),
         dly4_1 = dly4 - lag(dly4))

# split sample into two parts - estimation sample and prediction sample
fstQ <- 1960.00 # 1960Q1
lstQ <- 1978.75 # 1978Q4
tbl.wide.1 <- tbl.wide.all %>%
  filter(index <= as.yearqtr(lstQ))
```

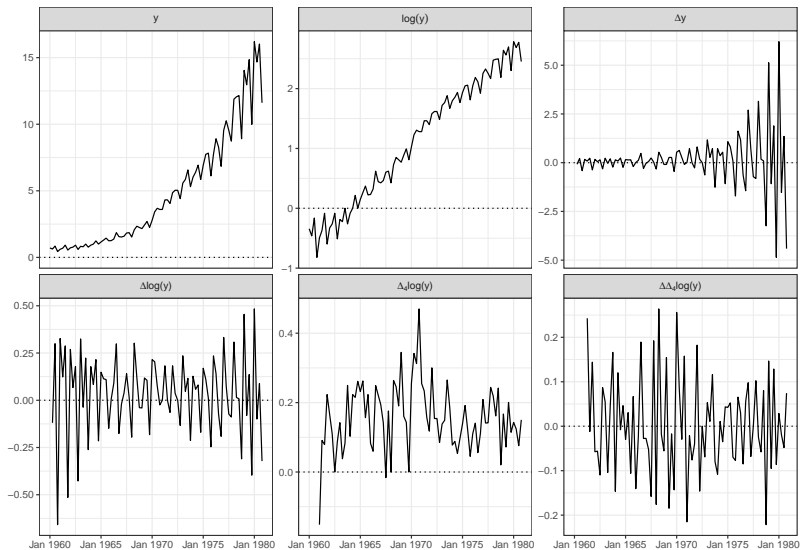
## Original and transformed data

```
# set default theme for ggplot2
theme_set(theme_bw())

# plot time series: levels, logs, differences
tbl.wide.all %>%
  gather(variable, value, -index) %>%
  mutate(variable.f = factor(variable, ordered = TRUE,
                             levels = c("y", "ly", "dy", "dly1", "dly4", "dly4_1"),
                             labels = c("y", "log(y)",
                                           expression(paste(Delta,"y")),
                                           expression(paste(Delta,"log(y)")),
                                           expression(paste(Delta[4],"log(y)")),
                                           expression(paste(Delta,Delta[4],"log(y)"))))) %>%

  ggplot(aes(x = index, y = value)) +
  geom_hline(aes(yintercept = 0), linetype = "dotted") +
  geom_line() +
  scale_x_yearmon() +
  labs(x = "", y = "") +
  facet_wrap(~variable.f, ncol = 3, scales = "free_y", labeller = label_parsed)
```

## Original and transformed data

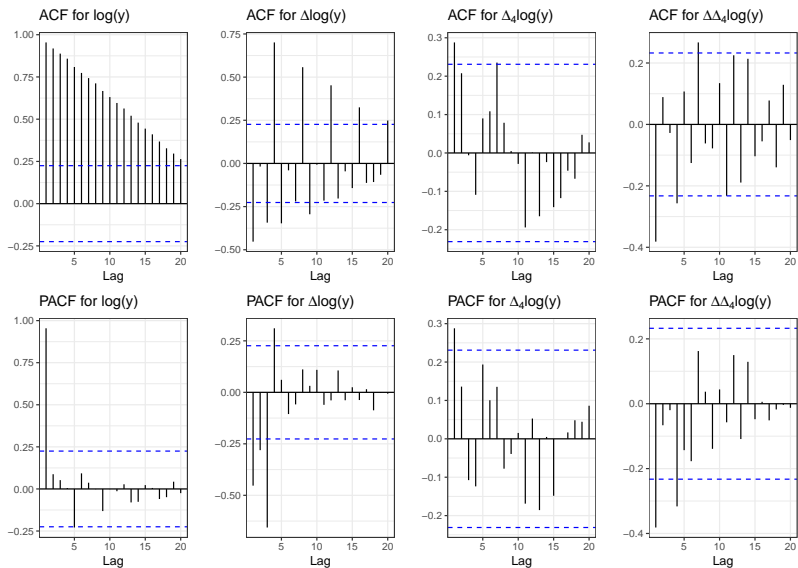


## ACFs and PACF

```
# plot ACF and PACF for 20 lags
maxlag <- 20
g1 <- ggAcf(tbl.wide.1$ly, lag.max = maxlag, ylab = "",
            main = expression(paste("ACF for log(y)")))
g2 <- ggAcf(tbl.wide.1$dly1, lag.max = maxlag, ylab = "",
            main = expression(paste("ACF for ", Delta, "log(y)")))
g3 <- ggAcf(tbl.wide.1$dly4, lag.max = maxlag, ylab="",
            main = expression(paste("ACF for ", Delta[4], "log(y)")))
g4 <- ggAcf(tbl.wide.1$dly4_1, lag.max = maxlag, ylab="",
            main = expression(paste("ACF for ", Delta, Delta[4], "log(y)")))
g5 <- ggPacf(tbl.wide.1$ly, lag.max = maxlag, ylab="",
            main = expression(paste("PACF for log(y)")))
g6 <- ggPacf(tbl.wide.1$dly1, lag.max = maxlag, ylab="",
            main = expression(paste("PACF for ", Delta, "log(y)")))
g7 <- ggPacf(tbl.wide.1$dly4, lag.max = maxlag, ylab="",
            main = expression(paste("PACF for ", Delta[4], "log(y)")))
g8 <- ggPacf(tbl.wide.1$dly4_1, lag.max = maxlag, ylab="",
            main = expression(paste("PACF for ", Delta,Delta[4], "log(y)")))

library(egg)
ggarrange(g1, g2, g3, g4, g5, g6, g7, g8, nrow = 2)
```

# ACFs and PACFs





## Estimate ARIMA(0, 1, 1)(0, 1, 1)<sub>4</sub> model

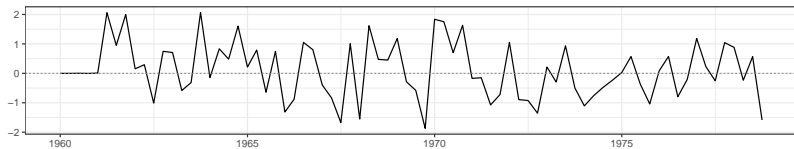
```
# estimate model
m1 <- tbl.wide.1 %>%
  tk_ts(select = ly, start = fstQ, frequency = 4) %>%
  Arima(order = c(0,1,1), seasonal = list(order = c(0,1,1), period = 4))
m1

## Series: .
## ARIMA(0,1,1)(0,1,1)[4]
##
## Coefficients:
##          ma1      sma1
##      -0.6559  -0.3492
## s.e.   0.1094   0.1104
##
## sigma^2 estimated as 0.008652: log likelihood=68.28
## AIC=-130.57  AICc=-130.21  BIC=-123.78
```

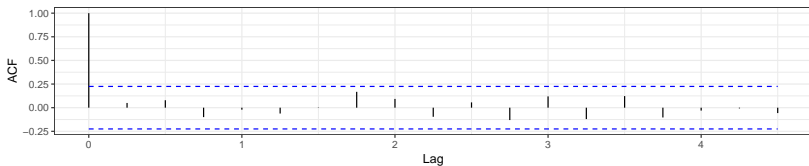
# Check Model Adequacy

```
ggtsdiag(m1, gof.lag=36)
```

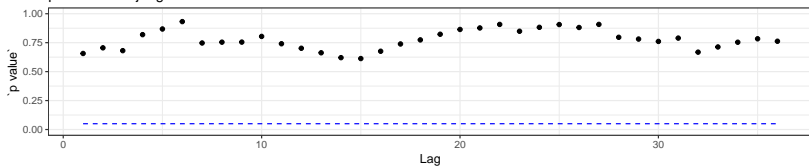
Standardized Residuals



ACF of Residuals



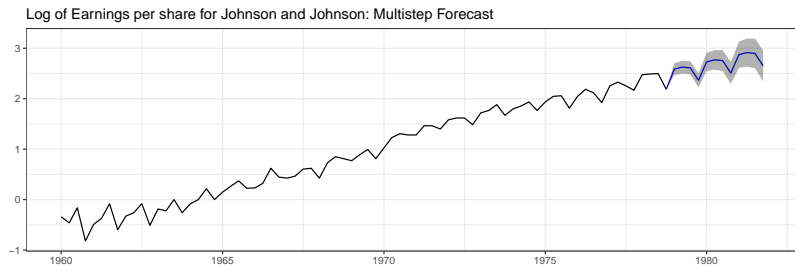
p values for Ljung-Box statistic



## Forecasts - Multistep

```
# construct 1-quarter to 12-quarters ahead forecasts
hmax <- 12
m1.f.1.to.hmax <- forecast(m1, hmax)

# plot 1-quarter to 12-quarters ahead forecasts - logs
autoplot(m1.f.1.to.hmax) +
  labs(x = "", y = "",
       title = "Log of Earnings per share for Johnson and Johnson: Multistep Forecast")
```



## Forecasts - Multistep

```
# actual data
tbl.2 <-
  tbl.wide.all %>%
  mutate(key = "actual",
         date = as.Date(index)) %>%
  select(date, key, y)

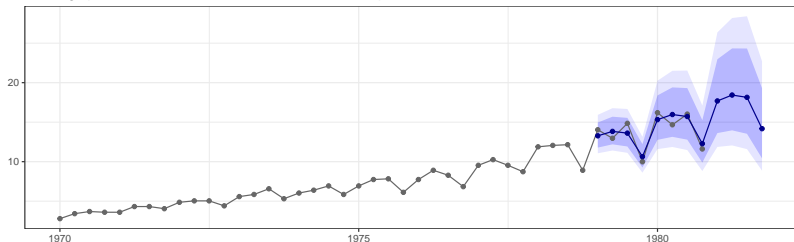
# extract the multistep forecasts, convert to levels
tbl.f.1.to.hmax <-
  m1.f.1.to.hmax %>%
  sw_sweep() %>%
  filter(key == "forecast") %>%
  mutate_at(vars(ly, lo.80, lo.95, hi.80, hi.95), funs(exp)) %>%
  mutate(date = as.Date(index)) %>%
  rename(y = ly) %>%
  select(date, key, y, lo.80, lo.95, hi.80, hi.95)

# forecast & actual data in a single tibble
tbl.f.1.to.hmax <- bind_rows(tbl.2, tbl.f.1.to.hmax)
```

## Forecasts - Multistep

```
# plot 1-quarter to 12-quarters ahead forecasts - levels
tbl.f.1.to.hmax %>%
  filter(date >= "1970-01-01") %>%
  ggplot(aes(x = date, y = y, col = key)) +
    geom_ribbon(aes(ymin = lo.95, ymax = hi.95), linetype = "blank", fill = "blue", alpha = 0.1) +
    geom_ribbon(aes(ymin = lo.80, ymax = hi.80), linetype = "blank", fill = "blue", alpha = 0.2) +
    geom_line() +
    geom_point() +
    scale_color_manual(values = c("gray40", "darkblue")) +
    labs(x = "", y = "",
         title = "Earnings per share for Johnson and Johnson: Multistep Forecast") +
    theme(legend.position = "none")
```

Earnings per share for Johnson and Johnson: Multistep Forecast



## Forecasts - Rolling Scheme

```
# window length for rolling SARIMA
window.length <- nrow(tbl.wide.1)

# create rolling SARIMA function with rollify from tibbletime package
roll_sarima <- rollify(~Arima(.x, order = c(0,1,1), seasonal = list(order = c(0,1,1), period = 4)),
                      window = window.length, unlist = FALSE)

# estimate rolling SARIMA model, create 1 step ahead forecasts
results <-
  tbl.wide.all %>%
  mutate(date = as.Date(index)) %>%
  as_tbl_time(index = date) %>%
  mutate(sarima.model = roll_sarima(ly)) %>%
  filter(!is.na(sarima.model)) %>%
  mutate(sarima.coefs = map(sarima.model, tidy, conf.int = TRUE),
         sarima.fcst = map(sarima.model, (. %>% forecast(1) %>% sw_sweep())))

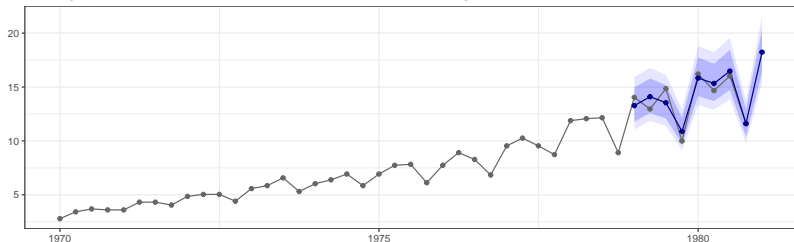
# extract the 1 period ahead rolling forecasts, convert to levels
m1.f.1.rol <-
  results %>%
  select(date, sarima.fcst) %>%
  unnest(sarima.fcst) %>%
  filter(key == "forecast") %>%
  mutate(date = date %m+% months(3)) %>%
  mutate_at(vars(value, lo.80, lo.95, hi.80, hi.95), funs(exp)) %>%
  rename(y = value) %>%
  select(date, key, y, lo.80, lo.95, hi.80, hi.95)

# forecast & actual data in a single tibble
tbl.f.1.rol <- bind_rows(tbl.2, m1.f.1.rol)
```

## Forecasts - Rolling Scheme

```
# plot 1-quarter ahead rolling forecasts - levels
tbl.f.1.rol %>%
  filter(date >= "1970-01-01") %>%
  ggplot(aes(x = date, y = y, col = key)) +
    geom_ribbon(aes(ymin = lo.95, ymax = hi.95), linetype = "blank", fill = "blue", alpha = 0.1) +
    geom_ribbon(aes(ymin = lo.80, ymax = hi.80), linetype = "blank", fill = "blue", alpha = 0.2) +
    geom_line() +
    geom_point() +
    scale_color_manual(values = c("gray40", "darkblue")) +
    labs(x = "", y = "",
         title = "Earnings per share for Johnson and Johnson: 1-Step Ahead Rolling Forecast") +
    theme(legend.position = "none")
```

Earnings per share for Johnson and Johnson: 1-Step Ahead Rolling Forecast



## Forecasts - Accuracy

```
# convert actual data in prediction sample into ts format
```

```
y.ts.2 <- tbl.wide.all %>%  
  filter(index > as.yearqtr(1stQ)) %>%  
  tk_ts(select = y, start = 1stQ+0.25, frequency = 4)
```

```
# evaluate accuracy of forecasts - multistep forecast - logs
```

```
accuracy(m1.f.1.to.hmax$mean, log(y.ts.2))
```

```
##           ME      RMSE      MAE      MPE      MAPE      ACF1 Theil's U  
## Test set -0.006368645 0.0640049 0.06092302 -0.3352225 2.347299 -0.8052326 0.2240991
```

```
# evaluate accuracy of forecasts - 1 step ahead rolling scheme forecast - logs
```

```
accuracy(log(m1.f.1.rol$y), log(y.ts.2))
```

```
##           ME      RMSE      MAE      MPE      MAPE      ACF1 Theil's U  
## Test set -0.008761921 0.06052803 0.05165746 -0.4089234 2.003344 -0.8127797 0.2103379
```

```
# evaluate accuracy of forecasts - multistep forecast - levels
```

```
accuracy(exp(m1.f.1.to.hmax$mean), y.ts.2)
```

```
##           ME      RMSE      MAE      MPE      MAPE      ACF1 Theil's U  
## Test set -0.0379115 0.8878123 0.8350954 -0.8424076 6.124417 -0.7756951 0.2255927
```

```
# evaluate accuracy of forecasts - 1 step ahead rolling scheme forecast - levels
```

```
accuracy(m1.f.1.rol$y, y.ts.2)
```

```
##           ME      RMSE      MAE      MPE      MAPE      ACF1 Theil's U  
## Test set -0.08727194 0.8036864 0.6987512 -1.06022 5.20146 -0.7713201 0.1990854
```